

META-ANALYSES OF THE EPIDEMIOLOGICAL EVIDENCE RELATING ETS TO LUNG CANCER AND HEART DISEASE

INTRODUCTION

Epidemiological evidence relating to two possible effects of ETS in never smokers is presented elsewhere, based on over 60 studies of lung cancer¹ and over 30 studies of heart disease.² In attempting to assess this overall evidence, the technique of meta-analysis is often used to combine the relative risk estimates from the individual studies. Results from a number of meta-analyses are presented in this document. However, before presenting these results, attention should be drawn to some features and limitations of the technique.

Meta-analysis was originally designed to combine evidence from randomised studies of similar design, to try to obtain a more accurate estimate of the possible effect of treatment. However, in recent years it has been increasingly used to combine evidence from epidemiological studies of quite widely varying design. It is important to realise that such a procedure may result in a combined relative risk estimate that has narrow confidence limits and therefore appears precise, but is in fact a quite inaccurate estimate of the true effect of treatment (if any).³

There are three major potential problems with meta-analysis.

The first, which also applies to the meta-analysis of randomised studies, lies in the possibility that the studies being analysed are unrepresentative of all those studies that actually exist. If, for example, editors are more likely to publish studies showing positive results than they are to publish studies showing negative results, and meta-analysis is based only on published evidence, such "publication bias" will lead to over-estimation of the true effect of treatment. Publication bias can result from various other sources also, and though various methods have been proposed for detecting and correcting for it, all have limitations.⁴

The second problem, specific to epidemiological studies, is that the results of the individual studies may be distorted by confounding and the other sources of bias that would be avoided in randomised studies. Some types of bias may be common to all (or most) of the studies. For example, it would be expected that studies would tend to find a higher risk of liver cirrhosis in heavy smokers than in nonsmokers, simply because heavy smokers are more likely to be heavy drinkers. Meta-analysis would not remove such a bias, and therefore may be misleading. A statistically significant result from a meta-analysis of epidemiological data does not necessarily establish that any underlying effect exists.^{5,6}

The final problem, which is particularly important in epidemiological studies, arises because of differences between the studies being combined. Subjects classified as exposed may, for example, have higher average exposure to the agent of interest in some studies than in others, so that, even in the absence of any bias, estimates of the true effect of exposure might not be expected to be the same in each study. Many other features of the study are also likely to add to the heterogeneity. These include aspects of study quality, such as the study design used and the accuracy of measurement of exposure and disease. They also include the nature of the population studied and where and when the study was conducted. If one is calculating a single combined estimate based on heterogeneous data - averaging apples and oranges in a sense - there is an obvious difficulty in interpreting the combined meta-analysis estimate.⁷

There are two main methods of conducting meta-analysis. One approach, **fixed-effects meta-analysis**, assumes that all the individual study estimates derive from a common mean, with their contribution to the overall estimate depending only on within-study variability, with large studies carrying more weight than do small studies. The alternative approach, **random-effects meta-analysis**, assumes that the individual study estimates derive from a distribution of effects, with the weighting of the individual estimates depending both on the within-study and the between-study variability. Both methods have their advantages and disadvantages.⁸ While the fixed-effects method totally ignores the not at all unlikely possibility of systematic heterogeneity between studies, the attempt to take heterogeneity into account by the random-effects method depends on an assumption about the nature of the distribution of effects over studies that is difficult to justify. Where heterogeneity exists, it is often held^{3,9,10} that more is to be gained by carefully examining its possible sources than by conducting random-effects meta-analysis. If, for example, the true situation is that the effect is consistently X in one group of studies and Y in another, then it would seem better to report the results as such than to attempt to produce a combined estimate.

While meta-analysis has its problems, it nevertheless remains of some value. One might regard the resulting estimate as an indicator of the approximate magnitude of the association to be explained, to be compared with estimates of the approximate magnitude of the various biases that might exist. In the context of ETS and lung cancer, for example, Hackshaw *et al*¹¹ have claimed that the major sources of bias are insufficient to explain the magnitude of association, a view which has been hotly contested.^{12,13}

Combined estimates of relative risk for the various indices of exposure, as estimated by both fixed- and random-effects meta-analysis¹⁴ are set out below.

LUNG CANCER

Combining the overall incidence for each of the four most commonly used indices of ETS exposure, the meta-analysis results show no association with childhood or social exposure, but a possible weak association with smoking by the spouse and in the workplace.

Index of ETS exposure	Estimates combined	Meta-analysis relative risk (95% confidence limits)	
		Fixed effects	Random-effects
Smoking by the husband	69	1.17 (1.10-1.23)	1.21 (1.12-1.31)
Smoking by the wife	26	1.15 (0.99-1.35)	1.15 (0.99-1.35)
Workplace exposure	36	1.24 (1.14-1.34)	1.23 (1.13-1.35)
Childhood exposure from any cohabitant	33	1.06 (0.98-1.15)	1.16 (1.00-1.35)*
Childhood exposure from the mother specifically	9	0.96 (0.77-1.20)	0.98 (0.77-1.25)
Social exposure	12	1.04 (0.92-1.17)	1.02 (0.80-1.28)

*Inflated by one study¹⁵ reporting an extremely high estimate of 12.0 (4.30-32.0) for women.

Subdividing the data on smoking by the husband produces the following:

Studies	Estimates combined	Meta-analysis relative risk (95% confidence limits)	
		Fixed effects	Random-effects
Estimates published in 1981-1989	25	1.38 (1.23-1.55)	1.38 (1.23-1.55)
Estimates published in 1990-1999	27	1.09 (1.01-1.17)	1.15 (1.02-1.30)
Estimates published in 2000 onwards	17	1.17 (1.04-1.31)	1.16 (1.03-1.32)

It is notable that studies published since 1989 show a statistically significantly lower relative risk than those published earlier.

The meta-analysis relative risks for smoking by the husband and for workplace ETS exposure are statistically significant. However, the associations are weak and may be explained by various sources of bias, including:

- **Confounding**

Many of the studies on ETS and lung cancer fail to consider diet, lifestyle, family medical history, socio-economic status and other factors known to differ between smoking and nonsmoking households.^{13,16,17}

- **Misclassification**

Some of the subjects classified as lifelong non-smokers will in fact be current or past smokers who have failed to report this.^{13,18,19}

- **Recall bias**

The studies rely on reported rather than objectively measured ETS exposure data, which may be affected by presence or knowledge of disease.¹²

- **Publication bias**

Studies finding a negative relationship of ETS exposure with lung cancer may be less likely to report their findings than those that find a positive relationship.⁴

When all the results are considered, and even when meta-analysis is applied, claims that the epidemiological data on ETS and lung cancer support an inference of causality¹¹ cannot be convincingly justified.^{12,13}

HEART DISEASE

The technique of meta-analysis has also been used to combine the results of epidemiological studies on ETS and ischaemic heart disease. The breakdown of results in relation to smoking by the spouse and to workplace ETS exposure can be summarised as follows:

Studies	Estimates combined	Meta-analysis relative risk (95% confidence limits)	
		Fixed-effects	Random-effects
Spouse ever smoked*	48	1.10 (1.07-1.12)	1.17 (1.10-1.24)
Spouse current smoker [†]	47	1.11 (1.08-1.14)	1.19 (1.12-1.27)
Workplace ETS exposure	18	1.12 (1.01-1.23)	1.12 (1.01-1.23)

* Using estimates for “spouse ever smoked” where a study also provides data for “spouse a current smoker”.

[†] Using estimates for “spouse a current smoker” where a study also provides data for “spouse ever smoked”.

There is substantial variation in meta-analysis estimates by period of publication and size of study. It is notable that the relative risks from the two largest US studies, published in 1995²⁰ and 2003²¹, are very close to 1.00 in each sex, and not statistically significant. These studies provide data on a total of over 20,000 heart disease cases, not far short of the total number in all the other studies combined.

While the overall adjusted relative risk estimates for spousal smoking are statistically significant, they are based on heterogeneous estimates which are substantially higher in small than in large studies. Many of the studies failed to control adequately for confounding or the various other sources of bias present in such epidemiological studies, with none adjusting for misclassification of smoking habits. Heart disease studies show no clearly significant relationship with workplace ETS exposure.

Again, claims that the epidemiological data for heart disease support an inference of causality^{22,23} cannot be convincingly justified.²⁴

ETS META-ANALYSES FOR LUNG CANCER AND HEART DISEASE

Exposure	Estimates included	Number of estimates	Relative risk (95% confidence limits)	
			Fixed-effects	Random-effects
LUNG CANCER				
Husband's smoking*	All	69	1.17 (1.10-1.23)	1.21 (1.12-1.31)
	N America	21	1.11 (1.01-1.22)	1.11 (1.01-1.22)
	Europe	16	1.21 (1.07-1.37)	1.23 (1.05-1.44)
	Asia	32	1.19 (1.10-1.29)	1.29 (1.14-1.46)
	1981-86	12	1.30 (1.11-1.52)	1.30 (1.09-1.55)
	1987-89	13	1.50 (1.26-1.78)	1.50 (1.26-1.78)
	1990-94	11	0.99 (0.89-1.09)	1.04 (0.84-1.28)
	1995-98	11	1.23 (1.08-1.40)	1.25 (1.07-1.47)
	1999-2001	11	1.28 (1.10-1.49)	1.28 (1.10-1.49)
	2002 onwards	11	1.08 (0.94-1.24)	1.08 (0.94-1.24)
Wife's smoking*	All	26	1.15 (0.99-1.35)	1.15 (0.99-1.35)
	N America	11	1.06 (0.82-1.37)	1.06 (0.82-1.37)
Husband or wife smoked*	All	88	1.17 (1.11-1.23)	1.21 (1.12-1.29)
	N America	30	1.11 (1.01-1.21)	1.11 (1.01-1.21)
	Europe	20	1.20 (1.07-1.34)	1.21 (1.05-1.40)
	Asia	38	1.20 (1.11-1.30)	1.29 (1.14-1.45)
	1981-86	18	1.32 (1.14-1.53)	1.32 (1.13-1.54)
	1987-89	15	1.53 (1.28-1.81)	1.53 (1.28-1.81)
	1990-94	12	0.99 (0.89-1.10)	1.05 (0.85-1.29)
	1995-98	16	1.21 (1.08-1.37)	1.22 (1.07-1.39)
	1999-2001	12	1.24 (1.08-1.42)	1.25 (1.08-1.43)
	2002 onwards	15	1.08 (0.95-1.23)	1.08 (0.94-1.23)
Workplace ETS exposure	All	36	1.24 (1.14-1.34)	1.23 (1.13-1.35)
Childhood ETS exposure				
From any cohabitant	All	33	1.06 (0.98-1.15)	1.16 (1.00-1.35)
From the mother specifically	All	9	0.96 (0.77-1.20)	0.98 (0.77-1.25)
ETS exposure in social settings	All	12	1.04 (0.92-1.17)	1.02 (0.80-1.28)
HEART DISEASE				
Spouse ever smoked**†	All	48	1.10 (1.07-1.12)	1.17 (1.10-1.24)
	1984-88	7	1.22 (1.04-1.44)	1.26 (1.02-1.55)
	1989-92	11	1.29 (1.17-1.43)	1.39 (1.18-1.64)
	1993-96	9	1.03 (0.99-1.07)	1.04 (0.99-1.08)
	1997-2000	7	1.32 (1.09-1.61)	1.37 (1.06-1.77)
	2001 onwards	14	1.14 (1.10-1.19)	1.15 (1.04-1.28)
	1-99 cases	6	1.60 (1.24-2.06)	1.61 (1.24-2.10)
	100-199	10	1.37 (1.11-1.69)	1.37 (1.11-1.69)
	200-999	19	1.23 (1.13-1.35)	1.26 (1.12-1.41)
	1000+	13	1.08 (1.05-1.11)	1.09 (1.01-1.16)
	Spouse current smoker	All	47	1.11 (1.08-1.14)
Workplace ETS exposure	All	18	1.12 (1.01-1.23)	1.12 (1.01-1.23)

Note: The individual study data used for the meta-analyses and more detailed meta-analyses are given elsewhere for both lung cancer^{1,25,26} and for heart disease.^{2,27,28} The individual study estimates used in the table above are those adjusted for as many potential confounding variables as the original authors have presented results for.

* Where spousal smoking data are not available, data for the nearest equivalent exposure index have been used.

† The studies by Palmer and Mannino are omitted because confidence limits are not available, and the study by Spencer is omitted as neither relative risk nor confidence limit is available.

This work was supported by the tobacco industry. The accuracy of the material presented is the responsibility of the authors alone.

References

Style BMJ

1. Lee PN, Forey BA, Hamling JS. *Epidemiological evidence on environmental tobacco smoke and lung cancer*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006G]
2. Lee PN, Forey BA, Hamling JS. *Epidemiological evidence on environmental tobacco smoke and heart disease*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006H]
3. Egger M, Schneider M, Davey Smith G. Spurious precision? Meta-analysis of observational studies. *BMJ* 1998;**316**:140-4.
4. Thornton A, Lee P. Publication bias in meta-analysis: its causes and consequences. *J Clin Epidemiol* 2000;**53**:207-16.
5. Shapiro S. Meta analysis/Shmeta-analysis. *Am J Epidemiol* 1994;**140**:771-8.
6. Petitti DB. Of babies and bathwater. *Am J Epidemiol* 1994;**140**:779-82.
7. Peto R. Why do we need systematic overviews of randomized trials? [Transcript of an oral presentation, modified by the editors]. *Stat Med* 1987;**6**:233-44.
8. Jones DR. Meta-analysis: Weighing the evidence. *Stat Med* 1995;**14**:137-49.
9. Blair A, Burg J, Foran J, Gibb H, Greenland S, Morris R, *et al*. Guidelines for application of meta-analysis in environmental epidemiology. *Regul Toxicol Pharmacol* 1995;**22**:189-97.
10. Thompson SG. Why sources of heterogeneity in meta-analysis should be investigated. *BMJ* 1994;**309**:1351-5.
11. Hackshaw AK, Law MR, Wald NJ. The accumulated evidence on lung cancer and environmental tobacco smoke. *BMJ* 1997;**315**:980-8.
12. Lee PN, Fry JS, Forey BA. Revisiting the association between environmental tobacco smoke exposure and lung cancer risk. V. Overall conclusions. *Indoor Built Environ* 2002;**11**:59-82.
13. Lee PN, Fry JS. *The relationship between lung cancer and ETS exposure: adjustment for the potential confounding effects of multiple risk factors and for misclassification of active smoking status. Updated analyses*. P.N. Lee Statistics and Computing Ltd.: 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006N.pdf and appendices in LEE2006N_APP.zip]
14. Fleiss JL, Gross AJ. Meta-analysis in epidemiology, with special reference to studies of the association between exposure to environmental tobacco smoke and lung cancer: a critique. *J Clin Epidemiol* 1991;**44**:127-39.
15. Rapiti E, Jindal SK, Gupta D, Boffetta P. Passive smoking and lung cancer in Chandigarh, India. *Lung Cancer* 1999;**23**:183-9.
16. Thornton A, Lee P, Fry J. Differences between smokers, ex-smokers, passive smokers and non-smokers. *J Clin Epidemiol* 1994;**47**:1143-62.

17. Fry JS, Lee PN. Revisiting the association between environmental tobacco smoke exposure and lung cancer risk. II. Adjustment for the potential confounding effects of fruit, vegetables, dietary fat and education. *Indoor Built Environ* 2001;**10**:20-39.
18. Lee PN, Forey BA. Misclassification of smoking habits as determined by cotinine or by repeated self-report - a summary of evidence from 42 studies. *J Smoking-Related Dis* 1995;**6**:109-29.
19. Lee PN, Forey BA, Fry JS. Revisiting the association between environmental tobacco smoke exposure and lung cancer risk. III. Adjustment for the biasing effect of misclassification of smoking habits. *Indoor Built Environ* 2001;**10**:384-98.
20. LeVois ME, Layard MW. Publication bias in the environmental tobacco smoke/coronary heart disease epidemiologic literature. *Regul Toxicol Pharmacol* 1995;**21**:184-91.
21. Enstrom JE, Kabat GC. Environmental tobacco smoke and tobacco related mortality in a prospective study of Californians, 1960-98. *BMJ* 2003;**326**:1057-61. Full version available at <http://bmj.com/cgi/content/full/326/7398/1057>
22. Glantz SA, Parmley WW. Passive smoking and heart disease. Mechanisms and risk. *JAMA* 1995;**273**:1047-53.
23. Law MR, Morris JK, Wald NJ. Environmental tobacco smoke exposure and ischaemic heart disease: an evaluation of the evidence. *BMJ* 1997;**315**:973-80.
24. Lee PN, Roe FJC. Environmental tobacco smoke exposure and heart disease: a critique of the claims of Glantz and Parmley. *Hum Ecol Risk Ass* 1999;**5**:171-218.
25. Lee PN, Forey BA, Hamling JS. *ETS and lung cancer meta-analyses*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006J]
26. Lee PN, Forey BA, Hamling JS. *Detailed meta-analysis on ETS and lung cancer*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006E]
27. Lee PN, Forey BA, Hamling JS. *ETS and heart disease meta-analyses*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006K]
28. Lee PN, Forey BA, Hamling JS. *Detailed meta-analysis on ETS and heart disease*. Sutton, Surrey: P N Lee Statistics and Computing Ltd; 2006. www.pnlee.co.uk/reflist.htm [Download LEE2006F]